

When optimal means optimal Finding optimal distinguishers from the mathematical theory of communication

Annelie Heuser, Olivier Rioul, Sylvain Guilley

Cryptarchi 2014

Motivation

questions raised by the community

What distinguishes known distinguishers in terms of distinctive features?

Given a side-channel context what is the best distinguisher among all known ones?



Motivation

questions raised by the community

What distinguishes known distinguishers in terms of distinctive features?

Given a side-channel context what is the best distinguisher among all known ones?

question we would like to answer

Given a side-channel scenario what is the best distinguisher among all possible ones?

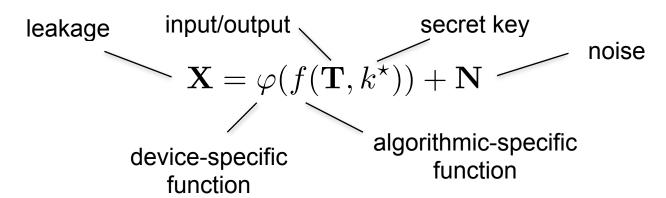


Outlook

- side-channel communication channel
- optimal distinguisher
 - known model
 - known model on a proportional scale
 - 1- bit models
 - partially known model
- empirical results
- what comes next!

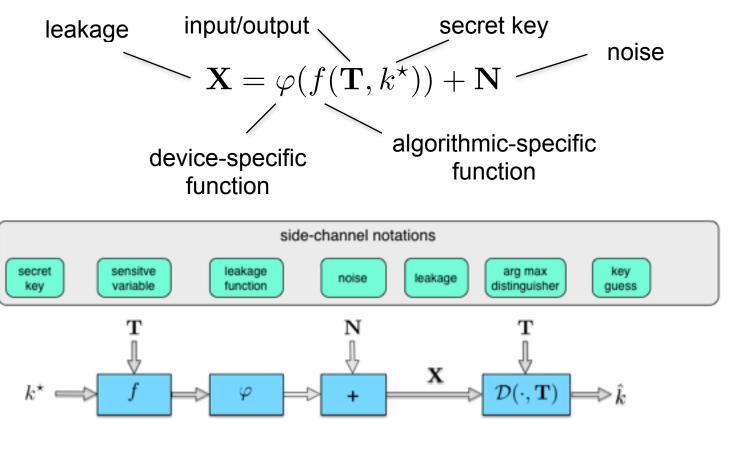


SCA as a communication channel





SCA as a communication channel







source

message

encoder

Optimal distinguishing rule

minimize the probability of error

$$\mathbb{P}_e = \mathbb{P}\{\hat{k} \neq k^*\}$$

Theorem (Optimal distinguishing rule) The optimal distinguishing rule is given by the maximum a posteriori probability (MAP) rule

$$\mathcal{D}(\mathbf{x}, \mathbf{t}) = \arg \max_{k} \left(\mathbb{P}\{k\} \cdot p(\mathbf{x}|\mathbf{t}, k) \right).$$

If the keys are assumed equiprobable, i.e. $\mathbb{P}\{k\} = 2^{-n}$, the equation reduces to the maximum likelihood distinguishing rule

$$\mathcal{D}(\mathbf{x}, \mathbf{t}) = \arg \max_{k} p(\mathbf{x} | \mathbf{t}, k)$$
.



Optimal distinguishing rule

minimize the probability of error

$$\mathbb{P}_e = \mathbb{P}\{\hat{k} \neq k^*\}$$

Theorem (Optimal distinguishing rule) The optimal distinguishing rule is given by the maximum a posteriori probability (MAP) rule

$$\mathcal{D}(\mathbf{x}, \mathbf{t}) = \arg \max_{k} \left(\mathbb{P}\{k\} \cdot p(\mathbf{x}|\mathbf{t}, k) \right) .$$

If the keys are assumed equiprobable, i.e. $\mathbb{P}\{k\} = 2^{-n}$, the equation reduces to the maximum likelihood distinguishing rule

$$\mathcal{D}(\mathbf{x}, \mathbf{t}) = \arg \max_{k} p(\mathbf{x}|\mathbf{t}, k)$$
.

Template attack [Chari+2002]



Optimal attack when the model is known

$$\mathbf{X} = \boldsymbol{\varphi}(\boldsymbol{f}(\mathbf{T}, k^{\star})) + \mathbf{N}$$

Proposition (Maximum likelihood) When f and φ are known to the attacker such that $\mathbf{Y}(k) = \varphi(f(k, \mathbf{T}))$, then the optimal decision becomes

$$\mathcal{D}(\mathbf{x}, \mathbf{t}) = \arg \max_{k} \left(\mathbb{P}\{k\} \cdot p(\mathbf{x}|\mathbf{y}(k)) \right) ,$$

and for equiprobable keys this reduces to

$$\mathcal{D}(\mathbf{x}, \mathbf{t}) = \arg \max_{k} \ p(\mathbf{x}|\mathbf{y}(k))$$
.



Optimal Attack when the model is known

Proposition When the leakage arises from $\mathbf{X} = \mathbf{Y}(k^*) + \mathbf{N}$, then

$$p(\mathbf{x}|\mathbf{y}(k)) = p_{\mathbf{N}}(\mathbf{x} - \mathbf{y}(k)) = \prod_{i=1}^{m} p_{N_i}(x_i - y_i(k)).$$

This expression depends only on the noise probability distribution $p_{\mathbf{N}}$.

- most publications considered Gaussian noise
- furthermore we investigate in uniform and Laplacian noise



Gaussian noise distribution

Theorem (Optimal expression for Gaussian noise) When the noise is zero mean Gaussian, $N \sim \mathcal{N}(0, \sigma^2)$, the optimal distinguishing rule is

$$\mathcal{D}_{opt}^{M,G}(\mathbf{x}, \mathbf{t}) = \arg\max_{k} \langle \mathbf{x} | \mathbf{y}(k) \rangle - \frac{1}{2} ||\mathbf{y}(k)||_{2}^{2}.$$



Gaussian noise distribution

Theorem (Optimal expression for Gaussian noise) When the noise is zero mean Gaussian, $N \sim \mathcal{N}(0, \sigma^2)$, the optimal distinguishing rule is

$$\mathcal{D}_{opt}^{M,G}(\mathbf{x}, \mathbf{t}) = \arg\max_{k} \langle \mathbf{x} | \mathbf{y}(k) \rangle \left(-\frac{1}{2} \| \mathbf{y}(k) \|_{2}^{2} \right)$$

- the optimal attack is independent on σ
- for large number of traces the last term becomes keyindependent but plays an important rule otherwise
- for large number of measurements the optimal distinguisher approximates to the covariance and the correlation
- but not with the absolute value!

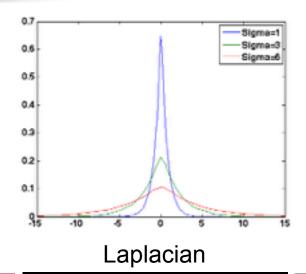


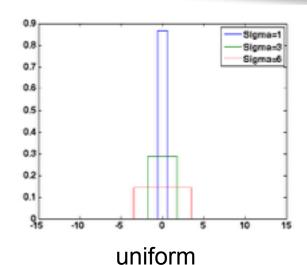
Uniform and Laplacian noise

Definition (Noise distributions) Let N be a zero-mean variable with variance σ^2 modeling the noise. Its distribution is:

• Uniform,
$$N \sim \mathcal{U}(0, \sigma^2)$$
 if $p_N(n) = \begin{cases} \frac{1}{2\sigma\sqrt{3}} & \text{for } n \in [-\sqrt{3}\sigma, \sqrt{3}\sigma] \\ 0 & \text{otherwise} \end{cases}$,

• Laplacian,
$$N \sim \mathcal{L}(0, \sigma^2)$$
 if $p_N(n) = \frac{1}{\sqrt{2}\sigma} e^{-\frac{|n|}{\sigma/\sqrt{2}}}$.







Uniform and Laplacian noise

Theorem (Optimal expression for uniform and Laplacian noises) When f and φ are known such that $Y(k) = \varphi(f(k,T))$, and the leakage arises from $X = Y(k^*) + N$ with $N \sim \mathcal{U}(0,\sigma^2)$ or $N \sim \mathcal{L}(0,\sigma^2)$, then the optimal distinguishing rule becomes

- Uniform noise distribution: $\mathcal{D}_{opt}^{M,U}(\mathbf{x},\mathbf{t}) = \arg\max_{k} -\|\mathbf{x} \mathbf{y}(k)\|_{\infty}$,
- Laplace noise distribution: $\mathcal{D}_{opt}^{M,L}(\mathbf{x},\mathbf{t}) = \arg\max_{k} \|\mathbf{x} \mathbf{y}(k)\|_{1}$.

- novel distinguishing rules
- cannot be approximated by correlation or covariance



Model known on a proportional scale

model only known on a proportional scale

$$X = aY(k^*) + b + N$$

where a and b are unknown and $a, b \in \mathbb{R}$

• one has to minimize $\|\mathbf{x} - a\mathbf{y}(k) - b\|_2$

Theorem (Correlation Power Analysis) Where N is zero-mean Gaussian, the optimal distinguishing rule becomes

$$\hat{k} = \arg\min_{k} \min_{a,b} \|\mathbf{x} - a\mathbf{y}(k) - b\|^2 ,$$

which is equivalent to maximizing the absolute value of the empirical Pearson's coefficient:

$$\hat{k} = \arg \max_{k} |\hat{\rho}(k)| = \frac{|\widehat{\text{Cov}}(\mathbf{x}, \mathbf{y}(k))|}{\sqrt{\widehat{\text{Var}}(\mathbf{x})\widehat{\text{Var}}(\mathbf{y}(k))}}.$$



11

Mono-bit leakage model

- w.l.o.g. $Y(k) = \pm 1$
- then $\|\mathbf{y}(k)\|_2^2$ is equal to the number of measurements

$$\mathcal{D}_{opt(1 \text{ bit})}^{M,G}(\mathbf{x}, \mathbf{t}) = \arg\max_{k} \langle \mathbf{x} | \mathbf{y}(k) \rangle = \arg\max_{k} \sum_{i | y_i(k) = 1} x_i - \sum_{i | y_i(k) = -1} x_i.$$

not equivalent to the difference-of-means test [Kocher+1999]

$$\mathcal{D}_{\mathrm{KJJ}}^{M,G}(\mathbf{x}, \mathbf{t}) = \arg\max_{k} \ \overline{\mathbf{x}_{+1}} - \overline{\mathbf{x}_{-1}}$$

nor to the t-test improvement [Coron+2000]

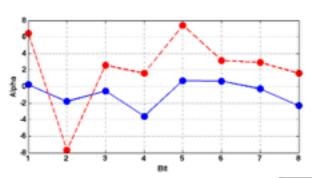


Model only partially known

leakage arising from a weighted sum of bits

$$X = \sum_{j=1}^{n} \alpha_j [f(T, k^*)]_j + N$$

- weights are unknown, epistemic noise is present
- assumption about the weights
 - unknown
 - normally distributed
 - fixed over one experiments





Model only partially known

Theorem (Optimal expression when the model is partially unknown)

Let $\mathbf{Y}_{\alpha}(k) = \sum_{j=1}^{n} \alpha_{j} [f(\mathbf{T}, k)]_{j}$ and $\mathbf{Y}_{j}(k) = [f(\mathbf{T}, k)]_{j}$. When assuming that the weights are independently deviating normally from the Hamming weight model, i.e., $\forall j \in [1, 8], \alpha_{j} \sim \mathcal{N}(1, \sigma_{\alpha}^{2})$, the optimal distinguishing rule is

$$\mathcal{D}_{opt}^{\alpha,G}(\mathbf{x},\mathbf{t}) = \arg\max_{k} \left(\gamma \langle \mathbf{x} | \mathbf{y}(k) \rangle + \mathbf{1} \right)^{t} \cdot \left(\gamma Z(k) + I \right)^{-1} \cdot \left(\gamma \langle \mathbf{x} | \mathbf{y}(k) \rangle + \mathbf{1} \right)$$
$$- \sigma_{\alpha}^{2} \ln \det(\gamma Z(k) + I) ,$$

where $\gamma = \frac{\sigma_{\alpha}^2}{\sigma^2}$ is the epistemic to stochastic noise ratio (ESNR), $\langle \mathbf{x} | \mathbf{y} \rangle$ is the vector with elements $(\langle \mathbf{x} | \mathbf{y}(k) \rangle)_j = \langle \mathbf{x} | \mathbf{y}_j(k) \rangle$, Z(k) is the $n \times n$ Gram matrix with entries $Z_{j,j'}(k) = \langle \mathbf{y}_j(k) | \mathbf{y}_{j'}(k) \rangle$, $\mathbf{1}$ is the all-one vector, and I is the identity matrix.

- if ESNR is small we recover the distinguisher when the model is known
- in contrast to linear regression the weights are not explicitly estimated



Empirical evaluation: known model

■ known model, only stochastic noise $X = \mathsf{HW}[\mathtt{Sbox}[T \oplus k^\star]] + N \ Y = \mathsf{HW}[\mathtt{Sbox}[T \oplus k]]$

compared distinguisher

$$\mathcal{D}_{opt}^{M,G}(\mathbf{x}, \mathbf{t}) = \arg\max_{k} \langle \mathbf{x} | \mathbf{y}(k) \rangle - \frac{1}{2} \| \mathbf{y}(k) \|_{2}^{2}, \qquad \text{(Euclidean norm)}$$

$$\mathcal{D}_{opt-s}^{M,G}(\mathbf{x}, \mathbf{t}) = \arg\max_{k} \langle \mathbf{x} | \mathbf{y}(k) \rangle, \qquad \text{(Scalar product)}$$

$$\mathcal{D}_{opt}^{M,L}(\mathbf{x}, \mathbf{t}) = \arg\max_{k} -\|\mathbf{x} - \mathbf{y}(k)\|_{1}, \qquad \text{(Manhattan norm)}$$

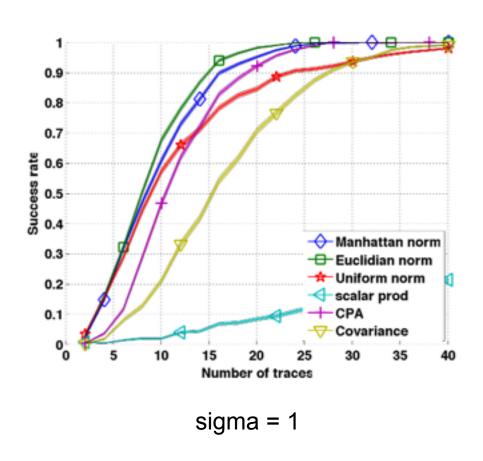
$$\mathcal{D}_{opt}^{M,U}(\mathbf{x}, \mathbf{t}) = \arg\max_{k} -\|\mathbf{x} - \mathbf{y}(k)\|_{\infty}, \qquad \text{(Uniform norm)}$$

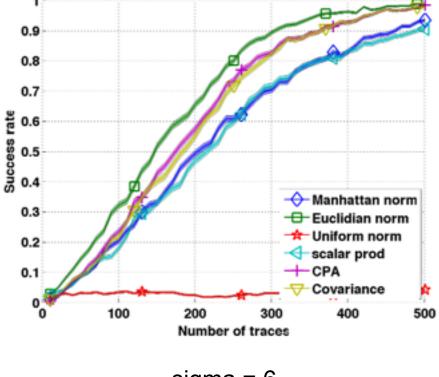
$$\mathcal{D}_{Cov}(\mathbf{x}, \mathbf{t}) = \arg\max_{k} |\langle \mathbf{x} - \overline{\mathbf{x}} | \mathbf{y}(k) \rangle|, \qquad \text{(Covariance)}$$

$$\mathcal{D}_{CPA}(\mathbf{x}, \mathbf{t}) = \arg\max_{k} \left| \frac{\langle \mathbf{x} - \overline{\mathbf{x}} | \mathbf{y}(k) \rangle}{\|\mathbf{x} - \overline{\mathbf{x}}\|_{2} \cdot \|\mathbf{y}(k) - \overline{\mathbf{y}(k)}\|_{2}} \right|. \qquad \text{(CPA)}$$



Gaussian noise

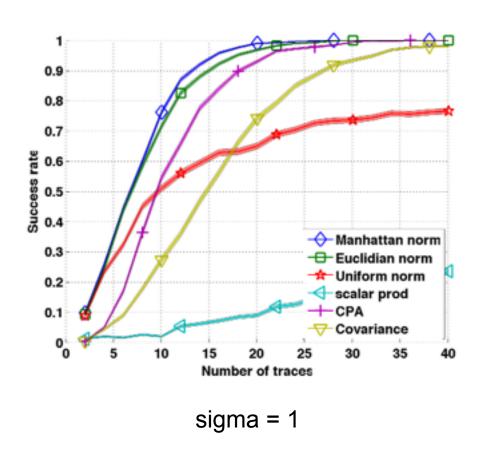


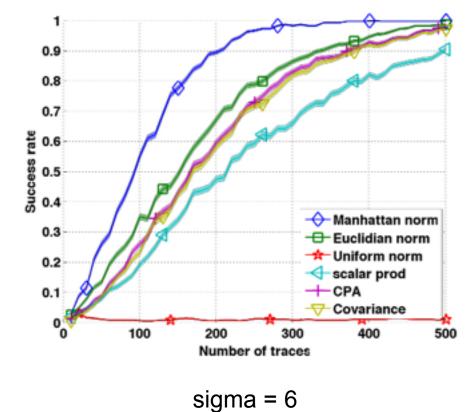






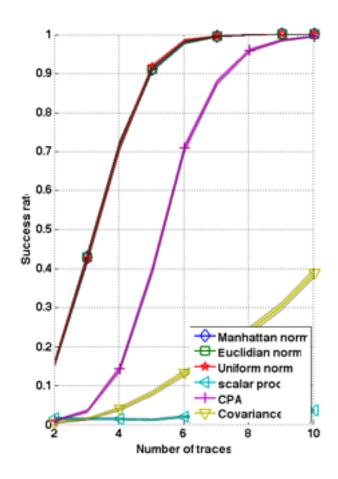
Laplacian noise

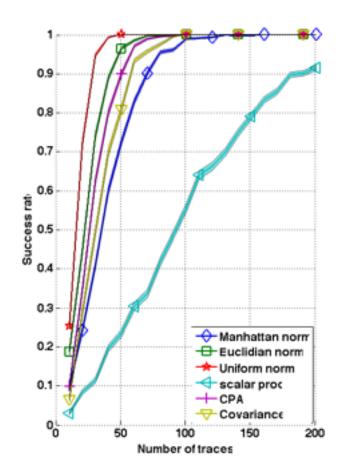






Uniform noise





$$sigma = 1$$



Gaussian noise: partially unknown model

stochastic scenario

$$Y_j = [\operatorname{Sbox}[T \oplus k]]_j \text{ for } j = 1, \dots, 8$$

$$X = \sum_{j=1}^8 \alpha_j Y_j(k^*) + N$$

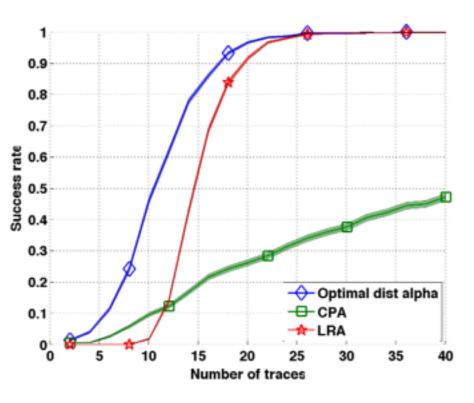
$$\alpha_j \sim \mathcal{N}(1, \sigma_\alpha)$$

 optimal distinguisher compared with linear regression attack (LRA)

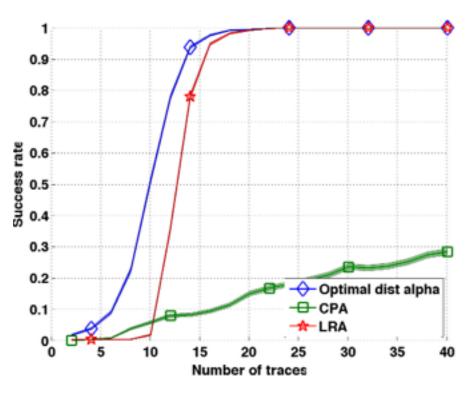
$$\mathcal{D}_{LRA}(\mathbf{x}, \mathbf{t}) = \arg \max_{k} \frac{\|\mathbf{x} - \mathbf{y}'(k) \cdot \boldsymbol{\beta}(k)\|_{2}^{2}}{\|\mathbf{x} - \overline{\mathbf{x}}\|_{2}^{2}},$$
$$\mathbf{y}'(k) = (\mathbf{1}, \mathbf{y}_{1}(k), \mathbf{y}_{2}(k), \dots, \mathbf{y}_{8}(k))$$



Gaussian noise: partially unknown model



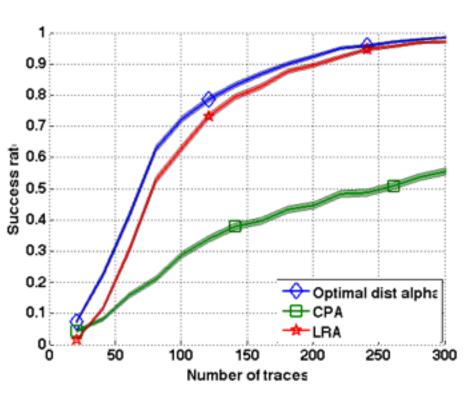
$$\sigma_{\alpha} = 2, \sigma = 1$$



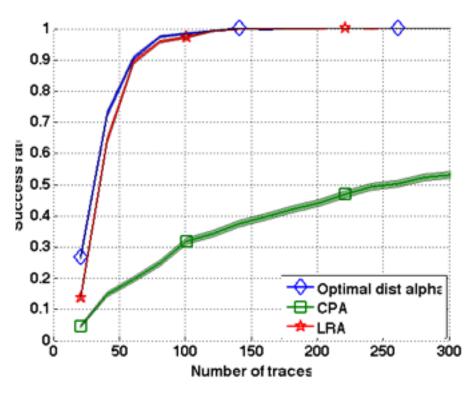
$$\sigma_{\alpha} = 4, \sigma = 1$$



Gaussian noise: partially unknown model



$$\sigma_{\alpha} = 2, \sigma = 6$$



$$\sigma_{\alpha} = 4, \sigma = 6$$



Conclusion

- transformed the problem of SCA into a communication theory problem to derive optimal distinguisher in a given context
- known leakage model:
 - Gaussian noise: optimal distinguisher close to CPA for low SNR
 - apart from Gaussian noise: optimal distinguishers differ from any known distinguisher
- partially unknown leakage model: optimal distinguisher performs better than LRA in the given context



Conclusion

- transformed the problem of SCA into a communication theory problem to derive optimal distinguisher in a given context
- known leakage model:
 - Gaussian noise: optimal distinguisher close to CPA for low SNR
 - apart from Gaussian noise: optimal distinguishers differ from any known distinguisher
- partially unknown leakage model: optimal distinguisher performs better than LRA in the given context

A mathematical study should prevail in side-channel analysis!



Future work

- quantify the gain in terms of numbers of traces required to break the key, in concrete setups (feasibility OK on DPA contest v4).
- preliminary step to determine the underlying scenario
- application to higher-order attack (under submission)

[Chari+2002] Suresh Chari, Josyula R. Rao, and Pankaj Rohatgi. Template Attacks. In CHES, volume 2523 of LNCS, pages 13–28. Springer, August 2002. San Francisco Bay(Redwood City), USA.

[Coron+2000] Jean-S´ebastien Coron, Paul C. Kocher, and David Naccache. Statistics and Secret Leakage. In Financial Cryptography, volume 1962 of Lecture Notes in Computer Science, pages 157–173. Springer, February 20-24 2000. Anguilla, British West Indies.

[Kocher+1999] Paul C. Kocher, Joshua Jaffe, and Benjamin Jun. Differential Power Analysis. In Proceedings of CRYPTO'99, volume 1666 of LNCS, pages 388–397. Springer-Verlag, 1999.

[Margard+2011] Stefan Mangard, Elisabeth Oswald, and Franc, ois-Xavier Standaert. One for All - All for One: Unifying Standard DPA Attacks. Information Security, IET, 5(2):100–111, 2011. ISSN: 1751-8709; Digital Object Identifier: 10.1049/iet-ifs.2010.0096.



Thank you!!

Questions?

to appear in CHES 2014, extended paper on eprint soon

Annelie Heuser is a Google European fellow in the field of privacy and is partially founded by this fellowship.

