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Template Attacks, Optimal Distinguishers & Perceived Information Metric

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Motivation

- Consolidate state-of-the-art about optimal distinguisher with a deeper look on the probabilities to estimate
- Perceived Information (PI) : information-theoretic metric quantifying the amount of leakage
- Show that PI is related to maximizing the success rate through the *Maximum a posteriori probability*
- Use the maximum likelihood to derive the template attack
- Experiments : If probabilities are known should they be used or estimated on-the-fly ?

Notations

- Secret key k^* deterministic but unknown
- m independent measurements $\mathbf{x} = (x_1, \dots, x_m)$ and independent and uniformly distributed inputs $\mathbf{t} = (t_1, \dots, t_m)$
- leakage model $\mathbf{y}(k) = \varphi(f(k, \mathbf{t}))$, where φ is a device specific leakage function and f maps the inputs to an intermediate algorithmic state.
- $\mathbf{x} = \mathbf{y}(k^*) + \mathbf{n}$ with independent noise \mathbf{n} .
- \mathbb{P} exact probability profiled device
- $\hat{\mathbb{P}}$ for an estimation offline (when profiling)
- $\tilde{\mathbb{P}}$ estimated online on-the-fly (when attacking)

Assumptions

The leakage model follows the

Markov condition

The leakage x depends on the secret key k only through the computed model $y(k)$. Thus, we have the Markov chain :

$$(k, t) \rightarrow y = \varphi(f(t, k)) \rightarrow x.$$

This assumption is related to the EIS assumption [SLP05].

Perceived information

Idea [RSVC⁺11]

- Metric quantifying degraded leakage models
- Generalization of mutual information
- Testing models against each other, e.g., from the true distribution against estimations

Ideal case

- the distribution \mathbb{P} is known
- PI is MI

$$MI(K; X, T) = H(K) + \sum_k \mathbb{P}(k) \sum_t \mathbb{P}(t) \sum_x \mathbb{P}(x|t, k) \log_2 \mathbb{P}(k|t, x)$$

Perceived information

Profiled case

- the distribution is known \mathbb{P}
- test a profiled model $\hat{\mathbb{P}}$ against \mathbb{P}

$$PI(K; X, T) = H(K) + \sum_k \mathbb{P}(k) \sum_t \mathbb{P}(t) \sum_x \mathbb{P}(x|t, k) \log_2 \hat{\mathbb{P}}(k|t, x)$$

Real case

- the distribution is unknown \mathbb{P}
- test a profiled model $\hat{\mathbb{P}}$ against an online estimated model $\tilde{\mathbb{P}}$

$$\hat{PI}(K; X, T) = H(K) + \sum_k \mathbb{P}(k) \sum_t \mathbb{P}(t) \sum_x \tilde{\mathbb{P}}(x|t, k) \log_2 \hat{\mathbb{P}}(k|t, x)$$

Maximum a posteriori probability

MAP

The optimal distinguishing rule is given by the *maximum a posteriori probability (MAP)* rule

$$\mathcal{D}(\mathbf{x}, \mathbf{t}) = \arg \max_k \mathbb{P}(k|\mathbf{x}, \mathbf{t}).$$

With the help of Bayes...

$$\mathbb{P}(k|\mathbf{x}, \mathbf{t}) = \frac{\mathbb{P}(\mathbf{x}|k, \mathbf{t}) \cdot \mathbb{P}(k)}{\mathbb{P}(\mathbf{x}|\mathbf{t})} = \frac{\mathbb{P}(\mathbf{x}|k, \mathbf{t}) \cdot \mathbb{P}(k)}{\sum_k \mathbb{P}(k)\mathbb{P}(\mathbf{x}|\mathbf{t}, k)}.$$

Relation between MAP and PI

Let \mathbb{P} be any distribution such that $\mathbb{P}(k|\mathbf{x}, \mathbf{t}) \propto \prod_{i=1}^m \mathbb{P}(k|x_i, t_i)$. We start by maximizing MAP :

$$\begin{aligned} \arg \max_k \hat{\mathbb{P}}(k|\mathbf{x}, \mathbf{t}) &= \arg \max_k \prod_{i=1}^m \hat{\mathbb{P}}(k|x_i, t_i) \\ &= \arg \max_k \prod_{x,t} \hat{\mathbb{P}}(k|x, t)^{m\tilde{\mathbb{P}}_k(x,t)}, \end{aligned}$$

where $\tilde{\mathbb{P}}_k(x, t) = \tilde{\mathbb{P}}(x, t|k)$ is the "counting" estimation (online) of x and t that depends on k . Now taking the \log_2 gives

$$= \arg \max_k \sum_{x,t} \tilde{\mathbb{P}}_k(x, t) \log_2 \hat{\mathbb{P}}(k|x, t)$$

Relation between MAP and PI (cont'd)

$$\begin{aligned} &= \arg \max_k \sum_{x,t} \tilde{\mathbb{P}}_k(x,t) \log_2 \hat{\mathbb{P}}(k|x,t) \\ &= \arg \max_k \sum_{x,t} \tilde{\mathbb{P}}(x,t|k) \log_2 \hat{\mathbb{P}}(k|x,t) \\ &= \arg \max_k \sum_t \tilde{\mathbb{P}}(t) \sum_x \tilde{\mathbb{P}}(x|t,k) \log_2 \hat{\mathbb{P}}(k|x,t) \end{aligned}$$

Taking the average over k and adding $H(K)$ gives $\hat{P}I(K; X, T)$

$$H(K) + \sum_k \mathbb{P}(k) \sum_t \tilde{\mathbb{P}}(t) \sum_x \tilde{\mathbb{P}}(x|t,k) \log_2 \hat{\mathbb{P}}(k|x,t).$$

(except $\tilde{\mathbb{P}}(t)$ vs. $\mathbb{P}(t)$)

Relation between MAP and PI (cont'd)

PI \Leftrightarrow MAP

$\hat{P}I$ is the expectation of the MAP over the keys.

Profiled case

If we have an infinite number of traces to estimate $\tilde{\mathbb{P}} \rightarrow \mathbb{P}$ then we recover $PI(K; X, T)$.

Ideal case

If we have an infinite number of traces to estimate $\tilde{\mathbb{P}} \rightarrow \mathbb{P}$ and $\hat{\mathbb{P}} \rightarrow \mathbb{P}$ then we recover $MI(K; X, T)$.

Maximum Likelihood Attack

Maximum Likelihood Attack

Assuming we have $y(k) = \varphi(f(t, k))$ that follows the Markov condition, then the optimal distinguishing rule is given by the maximum likelihood (ML) rule

$$\mathcal{D}(\mathbf{x}, \mathbf{t}) = \arg \max_k \mathbb{P}(\mathbf{x}|\mathbf{y}).$$

In practise...

- \mathbb{P} is most likely not known perfectly by the attacker
- either estimated offline by $\hat{\mathbb{P}}$
- or online on-the-fly $\tilde{\mathbb{P}}$

Maximum Likelihood Attack

Similarly, as in the previous derivation we have

$$\arg \max_k \mathbb{P}(\mathbf{x}|\mathbf{y}) = \prod_{i=1}^m \mathbb{P}(x_i|y_i) = \prod_{x,y} \mathbb{P}(x|y)^{m\tilde{\mathbb{P}}(x,y)}.$$

Taking the \log_2 gives us

$$\sum_{x,y} \tilde{\mathbb{P}}(x,y) \log_2 \mathbb{P}(x|y)$$

Now we add the cross entropy term that does not depend on a key guess k

$$- \sum_{x,y} \tilde{\mathbb{P}}(x,y) \log_2 \mathbb{P}(x)$$

Maximum Likelihood Attack

This results to

$$\arg \max_k \tilde{\mathbb{P}}(x, y) \log_2 \frac{\tilde{\mathbb{P}}(y|x)}{\tilde{\mathbb{P}}(y)}.$$

Profiled

\mathbb{P} is estimated offline $\hat{\mathbb{P}}$ on a training device

$$\arg \max_k \tilde{\mathbb{P}}(x, y) \log_2 \frac{\hat{\mathbb{P}}(y|x)}{\hat{\mathbb{P}}(y)},$$

which is the *template attack*.

Maximum Likelihood Attack

Non-Profiled

\mathbb{P} is estimated online $\tilde{\mathbb{P}}$ on a the device under attack

$$\arg \max_k \tilde{\mathbb{P}}(x, y) \log_2 \frac{\tilde{\mathbb{P}}(y|x)}{\tilde{\mathbb{P}}(y)},$$

which gives the Mutual Information Analysis [GBTP08].



Believing or seeing ?

Should probabilities considered as precise as possible ?

- Many recent works (e.g., [VCS09]) showed that using kernel estimation is more efficient than using histograms
- Accordingly, if $\mathbb{P}(Y)$ is known, should it be used instead of $\tilde{P}(Y)$ and $\hat{P}(Y)$?

Believing or seeing ?

Simple scenario

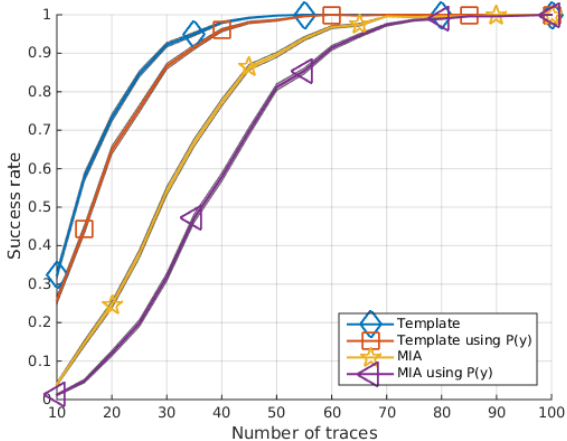
$$X = Y(k^*) + N,$$
$$Y(k) = HW(Sbox(T \oplus k))$$

As Y follows a binomial distribution with parameters $(n, 1/2)$, we have

$$\mathbb{P}(Y) = \{1/256, 8/256, 28/256, 56/256, 28/256, 8/256, 1/256\}.$$

- Template attack : replace $\tilde{\mathbb{P}}(Y)$ and $\hat{\mathbb{P}}(Y)$ by $\mathbb{P}(Y)$
- MIA : replace : $\tilde{\mathbb{P}}(Y)$ by $\mathbb{P}(Y)$

Believing or seeing ?



Conclusion

- $\hat{P}I(K; X, T)$ is the expectation of the MAP over the keys
- Maximum likelihood to recover
 - template attack when probabilities are estimated offline ($\hat{\mathbb{P}}$)
 - MIA when probabilities are estimated online on-the-fly ($\tilde{\mathbb{P}}$)
- In the attack phase : probabilities should be estimated instead of using the true distributions

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